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Regional Fertility Differences in India

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ABSTRACT

While theoretical literature distinguishes between factors that affect individual preferences regarding fertility and their ability to achieve these preferences, empirical literature often tends to conflate the two by focusing on completed family size. This chapter uses unique longitudinal data for India to distinguish between factors that affect fertility preferences, and those that affect ability to implement these preferences. India, with its tremendous regional heterogeneity in socioeconomic conditions as well as service delivery systems, offers a unique laboratory for this analysis. The results show that while socioeconomic characteristics of individuals account for substantial proportion of regional differences in fertility preferences, they only account for a small proportion of regional differences in unintended births. This suggests that unobserved factors, potentially those associated with regional health systems, have a far greater role in explaining underlying differences in unintended births than in explaining fertility preferences.

1. Introduction

Theoretical literature on fertility tends to differentiate between three sets of processes. First, individuals take into account their own social, and economic circumstances to develop a mental map of how many children (if any) they would like to have. Second, they negotiate with significant others in their lives and begin to develop a life plan for crystalizing their preferences, and finally they negotiate on the use of family planning services to find ways of implementing these preferences (Easterlin 1978, 1983; Easterlin and Crimmins 1985; Bulatao and Lee 1983, Hirschman 1994, Bongaarts 1978; Schoen et al. 1999).

Empirical literature has failed to keep up with this theoretical sophistication, resulting in strident debates in the field regarding the importance of family planning service delivery (see debates surrounding Pritchett 1994), or the role of innovation and diffusion vis-à-vis development, in shaping fertility outcomes (Cleland and Wilson, 1987). Part of the problem arises from the fact that much of the literature is based on cross-sectional data from large survey programs such as the World Fertility Survey, or Demographic and Health Surveys where we observe the ultimate culmination of all of these processes into achieved fertility, but do not have a step-by-step glimpse into how these factors play out in women's reproductive lives.

In this chapter we try to fill this niche by using unique longitudinal data for India to distinguish between factors that shape reproductive preferences from those that enable women to carry out their preferences. India provides an ideal laboratory to study the interplay between individual choices, and health services due to its geographic, economic, and cultural diversity overlaid with differences in state capacity and effectiveness.

Indian constitution divides up various functions of governance between the central, and the state governments; and health is under the portfolio of the state governments. While central government provides funding for many health, and family welfare programs including delivering contraception, and maternal and child health, ultimate administrative responsibilities lie with the states. State capacity varies substantially across states with some states carrying out their responsibilities with relative efficiency while others are unable to efficiently deliver services (Dreze and Sen, 2013).

Regional variation in fertility in India is well recognized; while fertility in many states is well below replacement level, in several large states it is still above replacement level, resulting in tremendous heterogeneity. In this chapter, we examine regional heterogeneity in fertility preferences, as well as women's ability to carry out these preferences to identify the extent to which this heterogeneity may be a function of individual characteristics such as education, and intra-household processes vis-à-vis that of a deeper systemic nature. An understanding of regional differences in fertility in India has tremendous policy significance. India's ability to achieve demographic transition rests on the ability of large states in North Central India such as Bihar, Uttar Pradesh and Madhya Pradesh to achieve fertility decline. Moreover, Indian government is in a process of changing the formulae for revenue sharing between in the center and the states with increased weightage being given to population share of individual states. This has led to complaints from the Southern states with claims that they are being punished for achieving demographic targets while the laggards are benefiting from their numerical superiority. A better understanding of the extent to which different parts of fertility are associated with state

performance vis-à-vis differences in the characteristics of the people who reside in these states may have important policy implications.

Using longitudinal data for the 19,132 women interviewed in both 2005 and 2012, we show that:

1. Regional differences in desired fertility as well as probability of having unplanned births are vast. In the so-called lagging states, desired fertility is 2.74 and probability of exceeding fertility preferences expressed in 2005 interviews in the subsequent seven years³ is 0.25; in contrast, for the states more advanced in demographic transition, desired fertility is 2.21 and probability of exceeding fertility preferences is 0.12.
2. A substantial proportion of the regional difference in desired fertility can be attributed to individual characteristics such as education and economic status. However, the likelihood of having an unplanned birth is only weakly associated with individual characteristics, leaving a far greater role for regional influences.

This chapter is organized into the following sections: section 2 includes discussion of fertility preferences, and ability to implement these preferences, sections 3 and 4 describe the sociocultural and demographic diversity of India, section 5 discusses the data used in the analyses, section 6 describes our conceptual framework for examining regional differences in fertility preferences, and behaviors, and section 7 focuses on statistical techniques, and construction of dependent variables. Section 8 describes the key independent variables used in this analysis, while sections 9, and 10 present descriptive and multivariate results respectively. Section 11 offers discussion of the results and concluding remarks.

2. Fertility Preferences Vs. Ability to Implement Preferences

In this chapter we focus on two dimensions of fertility – fertility preferences, and ability to implement these preferences. Both quantitative and qualitative studies find that fertility preferences are important in determining contraception usage, and fertility behavior (England et al., 2016; Hayford and Agadjanian, 2012; Moreau et al., 2013; Schoen et al., 1999; Yoo, Guzzo, and Hayford, 2014; Edin and Kefalas, 2011), however, this relationship is mediated by various factors. The translation of fertility preferences into behavior depends on context specific experiences faced by women (Dommaraju and Agadjanian, 2009; Agadjanian 2005; Johnson-Hanks, 2007).

The capacity of women to translate their fertility preferences into actual fertility is often constrained by a variety of factors – both within and outside the household. In particular, factors that affect crystallization of her preferences include ability to convince other household members, as well as knowledge about, and ability to obtain and effectively use contraception. Despite the various cross-sectional studies on fertility and family planning focusing on proximate determinants of contraceptive use such as socio-economic factors, and supply of family planning, longitudinal studies that examine the fertility preferences, and subsequent behavior are limited in number (Islam & Bairagi, 2003; Roy et al. 2008, Koenig et al. 2006; Kodzi et al. 2010; Vlassoff, 2012, Speizer et al. 2013, Kastor & Chatterjee 2018).

³ Sample size for this is 13,128 women who had children less than or equal to ideal family size in 2005.

Using longitudinal data, we first examine the correlates of women's fertility preferences measured by their response to questions regarding ideal family size. Then, we restrict our sample to women who are at, or below their ideal family size, and observe them seven years later to see how many exceed their preferred ideal, and factors that are associated with this phenomenon. We focus on these two processes to better understand regional diversity in fertility across Indian states.

3. Socio-Economic Diversity in India

Indian states are characterized by tremendous socioeconomic diversity. Table 1 shows some of the key characteristics of Indian states (aggregated from IHDS 2005 data). Because the IHDS samples are not designed to be representative at state level and in case of union territories, and north-eastern states are relatively small (Desai et al. 2009), we combine the 34 states and union territories into 22 groups where smaller states or Union Territories are combined with adjacent states. We present data on five indicators, average per capita income (in 2005 Rs.), percent of doctor assisted deliveries, percent households with access to flush toilets, percent households with access to electricity, and average number of consumer durables owned by a household.⁴

The primarily urban union territory of Delhi has the highest per capita income (Rs 15,000), followed by the Northeastern States (Rs 13,350), Kerala (Rs 9,987) and Himachal Pradesh (Rs 9,942). The poorest regions in the sample in 2005, that have the lowest per capita income are Orissa (Rs 3,450), followed by Bihar (Rs 3,530), Madhya Pradesh (Rs 4,125), and Uttar Pradesh (Rs 4,300). Household assets show the long-term economic standing of households. Together, assets and amenities such as electricity depict the overall quality of life. Column 5 in table 1 aggregates data from IHDS 2005 to the State level and we can see that Bihar, Chhattisgarh, Orissa, Jharkhand and Madhya Pradesh are the poorest States, and Delhi, Punjab, Haryana, and Kerala are the richest. From measures of per capita income/average number of assets owned, all of the poorest States are "Empowered Action Group (EAG)" States (greater explanation in section 4).

In terms of access to electricity, about 72% households in the IHDS 2005 sample had electricity in 2005 (column 4, Table 3). States where the lowest percentage households had access to any electricity were: Bihar (35%), Orissa (43%), Uttar Pradesh (45%), and West Bengal (53%). Only 23% households all over India (in the sample) had access to flush toilets. There was a wide state-wise variation in percentage households with access to flush toilets; while in Kerala and Delhi, it was more than 60%; in some of the poor States (Bihar, Orissa and, Chhattisgarh), it is below 10%. Finally, column (2), shows the percentage women in each state group who had a doctor assisted delivery (for their last birth). Physician assisted delivery is an indicator of the larger maternal care system in the state. In the IHDS 2005 sample, 42% women had a doctor assisted delivery, however there was a large state level variation in doctor assisted deliveries. While, it was the highest in the Southern States of Kerala (98%), Andhra Pradesh (82%), and

⁴ The IHDS asked questions on the goods households owned and quality of the household (on 30 items that include: a) ownership of goods such as television, refrigerator, chair/table etc.; and b) quality of housing (such as pucca wall, roof, electricity etc.).

Tamil Nadu (79%); it was the lowest in Uttar Pradesh (15%), Madhya Pradesh (16%), Uttarakhand (20%), and Chhattisgarh (21%). This is reflective of the poor maternal health care system in the EAG States compared to the non-EAG States.

Table 1. Contextual Characteristic's of the Key State Groups (in 2005)

| State | Per Capita Income (in Rs) (1) | Doctor Delivery for last birth (%) (2) | Household Access to Flush Toilets (%) (3) | Household Access to any Electricity (%) (4) | No. of Assets Owned (5) | Male Literacy (%) (6) | Female Literacy (%) (7) |
|-------------------|----------------------------------|---|--|--|----------------------------|--------------------------|----------------------------|
| All India | 5,999 | 42 | 23 | 72 | 11 | 81 | 60 |
| Jammu and Kashmir | 8,699 | 66 | 22 | 98 | 12 | 75 | 52 |
| Himachal Pradesh | 9,942 | 37 | 28 | 98 | 14 | 95 | 81 |
| Uttarakhand | 6,857 | 20 | 39 | 80 | 13 | 88 | 70 |
| Punjab | 9,125 | 47 | 43 | 97 | 18 | 85 | 72 |
| Haryana | 9,443 | 36 | 18 | 94 | 16 | 83 | 57 |
| Delhi | 15,000 | 62 | 64 | 99 | 19 | 92 | 78 |
| Uttar Pradesh | 4,300 | 15 | 13 | 45 | 10 | 76 | 48 |
| Bihar | 3,530 | 29 | 5 | 35 | 7 | 71 | 39 |
| Jharkhand | 4,833 | 33 | 13 | 61 | 9 | 78 | 51 |
| Rajasthan | 6,260 | 25 | 22 | 64 | 11 | 75 | 38 |
| Chhattisgarh | 5,306 | 21 | 7 | 68 | 8 | 76 | 50 |
| Madhya Pradesh | 4,125 | 16 | 24 | 76 | 9 | 78 | 49 |
| Northeast | 13,352 | 66 | 20 | 87 | 12 | 92 | 86 |
| Assam | 6,000 | 24 | 81 | 70 | 10 | 85 | 77 |
| West Bengal | 6,250 | 40 | 23 | 53 | 10 | 77 | 65 |
| Orissa | 3,450 | 36 | 5 | 43 | 8 | 81 | 60 |
| Gujarat | 6,300 | 57 | 40 | 88 | 14 | 86 | 63 |
| Maharashtra & Goa | 7,975 | 68 | 18 | 87 | 13 | 91 | 76 |
| Andhra Pradesh | 6,241 | 82 | 21 | 89 | 12 | 70 | 49 |
| Karnataka | 5,964 | 57 | 20 | 91 | 11 | 82 | 65 |
| Kerala | 9,987 | 98 | 67 | 90 | 16 | 98 | 97 |
| Tamil Nadu | 7,000 | 79 | 38 | 90 | 13 | 86 | 70 |

Source: IHDS 2005.

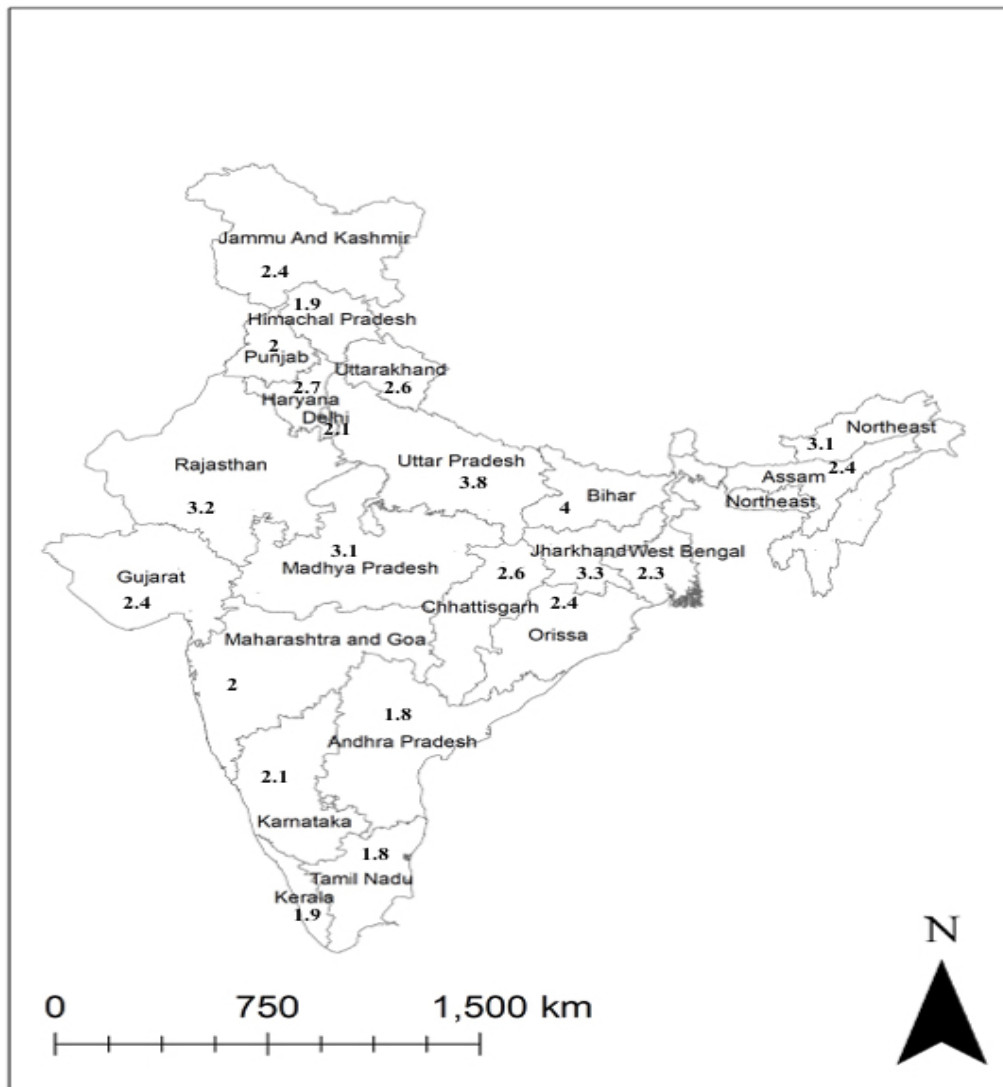
There is tremendous diversity between states in education. Column 6 and 7 of Table 1 shows proportion of men and women who were literate in 2005, with considerably higher level of literacy in Southern, and Western states of Kerala, Karnataka, Gujarat and Maharashtra, compared to states in Central India such as Uttar Pradesh, Bihar, and Madhya Pradesh.

4. Demographic Diversity in India

Socioeconomic diversity of India is also reflected in its demography. India is expected to surpass China by 2027, to become the most populous country in the World (United Nations Population Division, 2019). India's contribution to the yearly world population growth is higher than that of any other country (approximately 19 million out of 89 million). As in other developing countries, population growth in India is largely based in the poorest areas with high populations (PRB report on Census 2011).

The Indian sub-continent is marked by substantial regional diversity in culture, social systems and reproductive behavior (Dyson and Moore, 1983). While India has seen a marked decline in fertility over the decades, from a TFR of 6 children per women in the 1950s to 2.2 in 2017 (Population Bulletin 2015; RGI 2017), as Figure 1 documents, there is considerable variation in fertility rates across different state groups.

Figure 1. State-wise Total Fertility Rates (TFR) (15-49) using NFHS- III 2005-06



Source: Published report of National Family Health Survey (NFHS).

While a group of Northern States with high population, a larger prevalence of traditional norms and beliefs, lesser educational attainment, and less effective administration termed as “Empowered Action Group (EAG)” States (Bihar, Chhattisgarh, Jharkhand, Madhya Pradesh, Orissa, Rajasthan, Uttarakhand, and Uttar Pradesh), accounted for 46 percent of India’s population in 2011 and 53% of the growth in population, the Southern States comprised of 21% of India's total population in 2011 but contributed to just 15% of the national population growth from 2001 onwards.

Fertility rates are lower than the replacement level of 2.1 in 12 States, and union territories. Amongst the districts, 28% of all Indian districts (174 out of 621) have a fertility level below 2.1. It is anticipated that 12 districts all over India could achieve the lowest low fertility rates by 2020 (Guilmoto and Rajan, 2013). On the other hand, there are 72 districts that have fertility rates above 4. These districts are spread across the Northern and Northeastern region. The fall in fertility in these regions in the last decade has been lower than the national average.

5. India Human Development Survey (IHDS)

This chapter relies on data from the from the first nationally representative panel data from the India Human Development Survey (IHDS). IHDS is a nationally representative survey of 41,554 households (includes 215,754 individuals) that are spread across all the States and Union Territories of India (except for Andaman Nicobar and Lakshadweep), covering 384 districts, 1503 villages and 971 urban blocks. This survey was first conducted in 2005 and with the same households being re-interviewed with an 83% re-contact rate . Survey questions included information on income, consumption, education and morbidity experiences of various household members.

Additionally, ever-married women aged 15-49 were interviewed in Wave 1 (2004-5) about gender relations, fertility history, contraceptive use and number of ideal children. These women were again interviewed in 2011-12, presenting us with a longitudinal panel in which we can examine women’s ability to carry out their fertility preferences over a period of about 7 years. In the present study our sample is restricted to 20,464 ever married women aged 18-40 in 2005 who were interviewed in both rounds. We further restrict our focus to women aged 40 and below to ensure that early menopause, or infecundity does not bias our results. In addition sample size is limited (for descriptive results in Appendix Table 1, Tables 3,4), to non-missing data for each of the dependent, independent, and control variables. Specifically for the key dependent variables, descriptive analyses for a) ideal family size in 2005 is non-missing for 19,132 women; whereas the variable measuring b) unintended births between 2005 and 2012 is limited to 13,128 women (for whom the number of living children in 2005 was less than or equal to ideal family size in 2005). Finally, regression analyses presented in tables 5 and 6 is limited only to women who had at least one birth between January 2000, and their interview date in 2005; this is because data on maternal healthcare utilization is available only for these women. Thus, the sample size for the regression analysis in case of ideal family size as reported in 2005 is limited to 8,348 women. In addition, since the sample size for the binary outcome variable measuring at least one unintended birth between 2005 and 2012 is further limited to only those

who had children less than or equal to ideal family size in 2005, the sample size for the second regression analyses is limited to 5,903 women.⁵

6. Conceptual Framework for Explaining India's Demographic Diversity

India's demographic diversity, combined with longitudinal data, allows us to understand the role of socioeconomic development vis-à-vis state policy, and infrastructure in shaping fertility differentials between Indian states. Two separate processes may account for demographic diversity between states. As Table 1 indicates, north-central states are home to some of the poorest, and the least educated populations in India. They are also characterized by poorly functioning public health systems. Hence, we first examine the extent to which individual characteristics such as education, wealth, and intra-household relationships explain inter-state differences in fertility preferences. Thereafter, we examine the extent to which these same characteristics explain women's ability to implement their preferences, and avoid unwanted birth. We hypothesize that individual factors will play a greater role in explaining inter-state variation in fertility preferences, than in explaining unplanned and possibly unwanted births. The latter may be a function of state level public health systems, and family planning service delivery.

Research on fertility argues that both demand for children, and ability to avoid unwanted fertility (Bongaarts, 1994) form key components of total completed fertility; with socio-economic and ideational factors playing an important role in shaping demand for children, and contraceptive knowledge and availability shaping unwanted fertility. Unfortunately, empirical research often finds it difficult to distinguish neatly between the two. Demand for children, typically measured by questions regarding ideal family size, is affected by ex post rationalization (Lightbourne 1985; Bongaarts 1990, 2011; Westoff 1991; Bhushan and Hill 1996), i.e. ideal family size is adjusted in a way that it is closer to actual family size (even when desired family size is actually lower than actual family size). Consequently, measures of unwanted fertility are typically derived at the population level rather than at individual levels, making it difficult to examine factors driving individual women to have more children than they may consider ideal.

Using longitudinal data we are able to partially address this shortcoming, and try to distinguish between factors that may affect ideal family size, from those that affect unwanted fertility in subsequent years. In the present study, we focus on two indicators of fertility preferences and subsequent behavior.

6.1. *Ideal Family Size in 2005:*

International (Bulatao, 1981; Bulatao & Lee, 1983; Bankole, 1995) as well as India specific literature (Bongaarts, 2001; Dharmalingam et al., 2014) argues that demand for children plays an important role in shaping fertility. Ideal family size is often used as an indicator of demand for children in research focusing on India (for example Roy et al. 2008; Chatterjee and

⁵ In analyses not reported here, we removed the restriction women with a birth in the five years preceding the 2005 interview to expand the sample size. Our conclusions did not change. Hence, for parsimony only the final regression results are presented on the restricted sample.

Kastor 2018). We measure ideal family size using the following questions, administered to eligible women respondents at Wave 1 of the IHDS survey:

'If you could go back to the time you did not have any children and could choose the number of children to have in your life, how many would that be?'

This is a continuous variable that ranges from 0-12 in our data.

6.2. *Undesired births between 2005 and 2012:*

Unwanted fertility is far more difficult to measure due to ex-post rationalization and its aggregate measures have led to considerable debate (Pritchett, 1994; Bongaarts, 1994). Instead of relying on aggregate measures, we exploit longitudinal data to measure unplanned fertility. In order to construct the measure, we use the ideal family size variable in 2005; and limit our sample to those women who had living children fewer than, or equal to ideal family size in 2005. As the next step, we go on to see whether these women exceeded their ideal family size as reported in 2005 by 2012. If the total number of living children in 2012, is greater than the ideal family size in 2005, then the binary variable measuring unplanned births takes a value of 1; and the variable takes a value of 0 if the total number of living children as reported in 2012 is less than or equal to ideal family size in 2005.

7. Statistical Model:

Results presented in the paper are based on the hierarchical linear models which view individuals as being nested in within a state. Multilevel (random effects) models can be conveniently expressed as a system of equations at separate levels (Bryk and Raudenbush 1992; Goldstein 1995).

First, for the dependent variable measuring ideal family size in 2005, we consider a simple, two level random intercept model,

$$Y_{ij} = \beta + u_j^{(2)} + \epsilon_{ij}^{(1)} \text{-----} (1)$$

for the measurements i (level 1 observations)= 1, 2, ..., n_j ; and level 2 (state groups) $j=1, 2, \dots, M$. Y_{ij} is the outcome variable: ideal family size in 2005 for the i th person in the j th state, β an unknown fixed intercept, $u_j^{(2)}$ is a random intercept, and $\epsilon_{ij}^{(1)}$ is a level 1 error term. The assumption is that the errors are distributed normally with mean 0, and variance σ_1^2 ; random intercepts are assumed to be normally distributed with mean 0 and variance σ_2^2 , and to be independent of error terms (StataCorp LP, 2013). The interclass correlation coefficient (ICC) denoted by ρ for this model and is calculated as follows:

$$\rho = \text{Corr}(Y_{ij}, Y_{i'j}) = \sigma_2^2 / (\sigma_1^2 + \sigma_2^2) \text{.....} (2)$$

In model 1 we include only state level random effects; and in models 2 and 3 we add the other individual/household level variables of interest such as education, asset quintile, etc.; and variables indicating a woman's connection with the maternal and child health system (whether she obtained any ANC or PNC) in 2005.

Next, going on to the second dependent variable: a binary variable indicating whether a woman had an unintended birth between 2005 and 2012, we used mixed effects logistic regression. The most basic equation partitions the variance in unintended births across individuals, and states is:

$$\ln\left(\frac{P_{ij}}{1-P_{ij}}\right) = \beta_{0j}; \text{----- (3)}$$

$$\beta_{0j} = \delta_{00} + \mu_{0j}$$

Where p_{ij} reflects the probability of i th woman in j th state having an unintended birth. The logit of unintended birth is a function of a randomly varying state specific component β_{0j} . The state specific component is determined by the size of the random effects term μ_{0j} . The ICC can be calculated to estimate the importance of clustering in unintended births (Snijders and Bosker 2000) using the following formula:

$$ICC = \Gamma_0 / (\Gamma_0 + \Pi^2/3) \text{----- (4)}$$

Where Γ_0 is the estimated variance of the random effects term μ_{0j} , and Π is the quantity 3.14159. Using IHDS data, we find that without including any covariates, the ICC is 0.10. This suggests that for unintended births, about 10% of the variance lies between states and 90% is between women in a state. The goal is to reduce this unexplained between-state variance through inclusion of individual, and household factors that account for compositional differences between states.

The next step adds a series of individual and household level control variables, to take into account the compositional differences between districts. This equation includes:

$$\ln\left(\frac{P_{ij}}{1-P_{ij}}\right) = \beta_{0j} + \beta_{1j}X_{1i} \text{----- (5)}$$

Where X_{1i} represents woman specific characteristics such as birth cohort (age), parity, and education; and the other substantive variables of interest that are operationalized at an individual level (e.g. household wealth). In the final model 3 we add variables indicating a woman's contact with the maternal and child health system.

Change in the between-state variance reflects the importance of compositional differences across states. The above example describes our modeling strategy for the bivariate variable, unplanned fertility.

In summary, we estimate hierarchical regression models for two dependent variables: a) ideal family size in 2005; and b) women who have at least one unintended birth between 2005 and 2012. For each of these dependent variables, we run two sets of model: i) model 1: with only random intercepts for state groups; and ii) model 2: model 1+ all other individual and household characteristics (except ANC/ PNC use in 2005); and iii) model 3: model 2+ ANC and PNC use as reported in 2005. Model 3 is indicative of a woman's connection with the maternal and child health system, which in turn reflects availability of maternal healthcare facilities in the region. These hierarchical models that take into account geographical clustering at state level; and thus regional heterogeneity is taken into account (for e.g. these models account for the fact that women's ideal family size within states, are not independent, since women residing in different states could have different exposure to characteristics specific to that state). All analyses is done using Stata 15. The hierarchical models are estimated using Mixed command in STATA, and it includes random intercepts for the state of residence. Finally, for both sets of regressions we

calculate the Intra-Class Correlation Coefficient (ICC) in STATA, in order to estimate the importance of state-wise clustering in fertility intentions and behavior (Snijders and Bosker 2000; Desai and Wu 2010).

Since data on ANC/PNC is limited to women who had a child in the five years preceding the 2005 survey, sample for model 3 is smaller than that for model 2 (i) 8,348 compared to 19,132 for the dependent variable: ideal family size; and ii) 5,903 compared to 13,128 for the dependent variable measuring unintended births). In order to consistently analyze changes in ICC, we present all three models for the final smaller sample, but models estimated on the larger sample (i.e. 13,128 and 19,132 women respectively for the two dependent variables) show qualitatively similar results and do not affect the discussion in Section 10.⁶

8. Individual Level Determinants of Fertility

Our focus on examining inter-state differences in fertility relies on examining changes in inter-state variance by adding controls for known individual level correlates of fertility preferences and behaviors. While we control for a range of individual characteristics such as education and household income, we also control for somewhat more distal factors that have been hypothesized to shape both demand for children, and individuals' ability to implement their preferences:

8.1. Ideational Factors

Theories from sociology, psychology, cognitive sciences and communication studies propose various processes through which people form attitudes. Some of these processes would include socialization, processing social information, social learning and social influence. These theories focus on structural (such as an individual's position in a given network) and ideational (such as attitudes and beliefs of other members in a network) forces. Theories further stress on the importance of disseminating new information and putting them in practice amongst individuals belonging to various networks. Mead (1967) emphasized on the importance of role taking, and interaction with network members in the shaping of an individual's self (including his/her attitude). Various studies examine the impact of media on an individuals' development of their behavior and self-identity (e.g. Barber and Axinn, 2004; Bennett, 1975; Gamson et. al., 1992; Gamson and Modigliani, 1989).

The idea that smaller family size is preferable is often spread through various communication networks that may or may not be a part of family planning program initiatives. Under the thesis of ideational forces (Cleland and Wilson, 1987) empirical work may be categorized into research on a) diffusion; b) impact of mass media and c) social interaction of ideas about fertility and actual fertility behavior (Freedman, 1997). Bongaarts and Watkins (1996), Watkins (1992) and Watkins et al., (1995), studied diffusion through social interactions, in different types of natural groups. They emphasize the importance of social networks in the dissemination of ideas on fertility, and thus in shaping fertility preferences.

⁶ Tables available on request from authors.

Studies have suggested that newspaper, radio and television campaigns lead to a rise in available knowledge and communication on use of contraceptives and family planning issues, declines in fertility desires and rise in use of sterilization. These relationships have been found in contexts of various countries such as Iran, Brazil, Guatemala, Nigeria, Zambia, Columbia, Gambia etc. (review by Hornik and McAnany, 2001).

Technological innovation, and access to effective contraception has been an important determinant in bringing about declines in unintended pregnancy and fertility especially for those with comparatively bigger families (Ryder 1973; Westoff 1972; Westoff and Bankole 1996). Barber and Axinn, 2004 examine the role of mass media (as a means of social change) that impacts an individuals' behavior principally through ideational mechanisms. They use data for 1091 couples in the Chitwan Valley Family Study and find that exposure to mass media is associated with fertility behavior, inclination towards have smaller families, weaker son preferences and greater toleration of contraceptive use.

The first demographic transition in Europe, and the ongoing fertility decline in the developing countries can be somewhat attributable to the diffusion of new ideas and types of contraception (Bongaarts and Watkins, 1996). When individuals communicate on new ideas about fertility, family size, gender roles, information on experience and acceptance towards modern contraceptives their exposure to available information and their interaction with others impacts their attitudes towards their acceptability of high fertility and ways to restrict births (Montgomery and Casterline 1993,1996, Kohler 2001; Kohler et al. 2001; Buhler and Kohler 2004). Qualitative research conducted in Italy and Germany show that interpersonal communication is also seen as an important factor in determining fertility behavior (Bernadi 2003; Bernadi et al.2005).

The framework of the diffusion theory has been studied in the context of India particularly to falls in fertility in South India. In regions where fertility decline began earlier, the rate of diffusion was faster across social and cultural groups compared to regions where fertility decline begun later (Guilmoto and Rajan 2001). Dommaraju and Agadjanian (2009) find that the rate of diffusion in the Southern States was faster compared to the Northern States where it was almost absent. Appel et. al (2002) found that in southern India diffusion of low fertility took place across caste, religious and economic groups.

While a host of factors have been identified to promote diffusion of ideas regarding importance of family limitation and knowledge about contraceptive use (see National Research Council 1999), we focus on two key determinants. The first set of factors relate to the exposure of women to mass-media and the second relates to the connection of households with formal institutions. Role of newspapers, television and radios is often seen as key force promoting ideational change (Faria and Potter 1999). The second set of factors identifies connections to formal institutions. Studies (e.g. Basu and Sunder 1988) also document that social networks, and connections to individuals who travel in the larger world and may be exposed to what has come to be known as “developmental idealism” (Thornton et al. 2015), a complex of ideas regarding value of delayed marriage, smaller families and use of contraception.

Following is a description of how we operationalize these variables:

Exposure of women in the household to mass media: The IHDS survey asks whether women in the household were exposed to television, radio and newspaper. The response to this question was, none, sometimes, and regularly. We recode this to, no exposure, and at least some exposure (this combines the categories sometimes and regularly). Next, we recode the missing values on this question, to no exposure to any form of mass media. The index for exposure of women in the household to mass media is an additive index (ranging from 0-3), that takes a value of 0 if women in the household have no exposure to any form of mass media, and takes a value 3 if women in the household have at least some exposure to all 3 forms of mass media. The missing responses were recoded to 0 (indicating no exposure to mass media). We create dummy variables for each category of the index and take no exposure to mass media as the reference group.

Number of formal social networks: The household questionnaire in 2005 asked: ‘Do you or any members of your household have personal acquaintance with someone who works in any of the following occupation. a) Medical profession, b) Schools, and c) Government services.’ Each of these had a binary no (0), yes (1) response. We construct an additive index that counts the total number of formal social networks, and ranges from 0 (no formal social networks) to 3 (all three formal social networks). The missing responses were recoded to 0. We create dummy variables for each category of this index, and take no formal social networks as the reference group. A limitation of this variable is that we only look at whether the family member knows anyone in the formal sector, but does not capture the type of the network, this is because for a large proportion of our sample, data on type of network is missing, and we did not want to lose those cases.

8.2. Intra-Household Bargaining Power

Research and public policies assume that women’s ability to implement their fertility preferences is dependent upon their ability to obtain cooperation from other family members. Past studies have found that women’s empowerment influences her fertility intentions, and her ability to implement her intentions into actual behavior (for e.g. Upadhyay and Karasek 2012; Kishor and Subaiya 2008, Balk 1994, Hindin 2000, Upadhyay et al. 2014). In an extensive review article based on 60 studies examining the relationship between women’s empowerment and fertility, Upadhyay et al. 2014, find that the most commonly used measure of women’s empowerment was women’s role in household decision making (used in 37 of the 60 articles). In the present study, along the lines of assumptions by Upadhyay and Karasek 2012, we assume that, as a woman is more empowered, she will have higher aspirations for herself and her children, which would lead to her lowering the ideal number of children she wants, and the actual number of children she has, so that she has more resources to spend on herself and her children. We include markers of women’s decision-making authority in the household as a marker of gender empowerment.

Household Decision Making Ability: In the IHDS survey, eligible women in 2005 were asked who (respondent, husband, senior male, senior female, other, or no one) had the most say in the following household decisions: a) what to cook on a daily basis, b) whether to buy an expensive item such as television or refrigerator) how many children to have; if they have children: d) what to do if a child falls sick, e) whom should ones’ child marry. The response to each of these questions takes a value of 1 if the respondent (the eligible woman) has the most say, and takes a

value of 0 otherwise. We recode the missing values on this question, to 0. In order to construct this variable we look into all decisions except the decision on what to cook on a daily basis, since this is usually a decision that is taken by women. We construct an additive index for the number of decisions in which the respondent has the most say. This ranges from 0 (indicating that the woman doesn't have the most say in any of the household decisions), to 4 (indicating that the woman has the most say in all four decisions). We create dummy variables for each category of this index, and take most say in no decision as the reference group.

8.3. Contact with Health Systems

The next two variables measure a woman's contact with the maternal and child health care system, and is limited only to those who had a birth between January 2000, and the time when she was interviewed in 2005, and had valid data on maternal healthcare utilization for their last birth. If women have availed antenatal and postnatal check-ups in the past, it would mean that they have had a connection with reproductive health workers in the recent past.

Antenatal checkup (ANC): This is a binary variable measuring whether a woman obtained any antenatal checkup during her last pregnancy (as reported in 2005). It takes a value '1' if a woman obtained at least one antenatal check-up during her last pregnancy, and takes a value of '0' if a woman had obtained no antenatal checkup during her last pregnancy.

Postnatal checkup (PNC) index: This index takes a value ranging from 0 to 2. This variable takes a value of 0 if a woman and her child obtained no postnatal check-up, a value of 1 if at least one of them obtained a postnatal check-up, but more than 24 hours after birth but within 2 months of birth, and takes a value of 2 if at least one of them obtained a postnatal checkup within 24 hours of birth as per WHO recommendations (WHO 2014).

8.4. Socio-economic Individual and Household Characteristics

In addition to the more distal factors influencing fertility behaviors such as social networks and exposure to mass media, we also control for some of the key socio-economic household and individual level characteristics that have been shown in the past literature to affect fertility. These include:

Education, parity, and age: woman's education level (dummy variables are constructed for six categories namely: illiterate, some primary education, primary complete, secondary complete, higher secondary complete, and college or higher); number of living children in 2005 (parity in 2005); and age group (dummy variables are constructed for 5 age categories: 18-20, 21-25, 26-30, 31-35, and 36-40 respectively).

Household Wealth: Household asset quintile (dummy variables are constructed for 5 categories, ranging from the poorest to the richest households);

Caste and Religion: Caste and religion are frequently identified as primary axes along which social intercourse as well as social stratification takes place (Desai and Dubey, 2011; Desai et al. 2010). A salient characteristic of the Hindu religion is the division into castes. This classification is hierarchal in nature, and was historically based on occupation. Under article 341 of the Indian constitution, the lowest ranking castes are now part of an official list or schedule, and are referred to as Scheduled Castes (SC). In a similar manner, all indigenous tribes of India

are part of an official schedule under article 342 of the Indian constitution, and are referred to as Scheduled Tribes (ST). Historically both groups have lived on the margins of the mainstream Indian society. Scheduled Tribes, often live in isolated regions where employment and occupational opportunities as well as health services are limited. While Scheduled Castes are not geographically isolated, they also suffer from exclusion, particularly in access to health services (Sabharwal et al. 2014).

Hence we control for caste and religion of individuals by including indicators for caste group (dummy variables constructed for 4 categories namely, Forward Caste groups (such as Brahmins, Kayasthas, Kshatriyas, etc.); and lower caste groups: Scheduled Castes (SC), Scheduled Tribes (ST), and Other Backward Classes (OBC)); and indicator for religious group (dummy variables are created for each of the following categories: Hindu, Muslim and Other Religions).

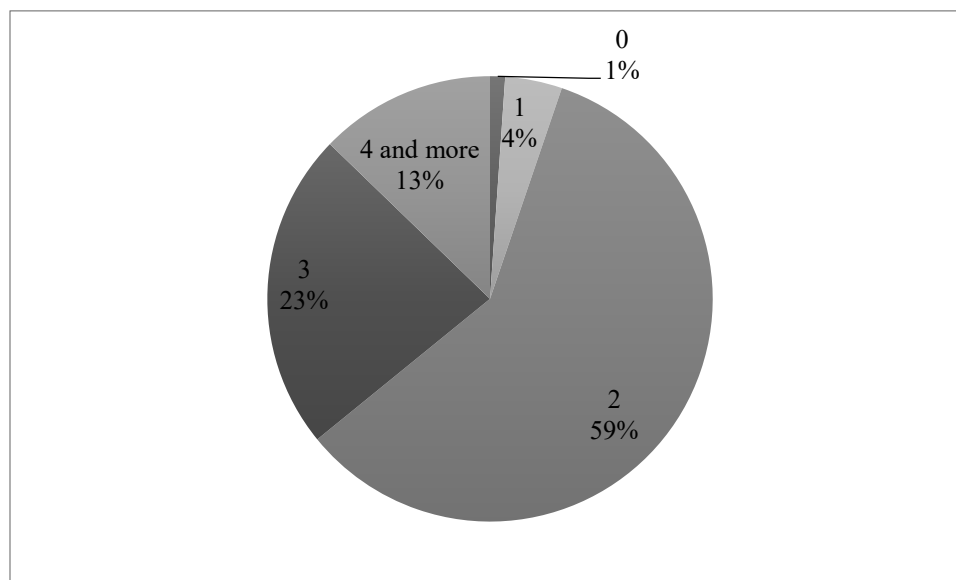
While an increase in education is expected to be associated with lower fertility (for e.g. Jejeebhoy, 1995; Murthi et. al, 1995); some studies in the Indian context have shown that Muslim women are likely to have higher fertility and lower contraception use compared to Hindu women (for e.g. Dharmalingam and Morgan 2004; Kulkarni and Algarajan, 2005). On the other hand the hypothesized relationship between caste and fertility is not very clear (Malhotra et al. 1995). While on one hand people belonging to lower castes are poorer, and people belonging to poorer families will be expected to have more children in order to obtain an additional source of labor income; on the other hand literature on social stratification indicates that 'status' concerns might lead to higher fertility amongst upper caste groups (Malhotra et al. 1995; Miller 1981).

Area of residence: This is a variable with 4 different categories namely: metro cities, other urban areas, more developed villages and less developed villages. There is a lot of variation in terms of urban and rural areas in India. While some large Indian cities have global influence, some other urban areas are hardly different from bigger villages (Desai et al. 2009). Therefore, the IHDS classifies urban regions into: a) metro cities (Mumbai, Kolkata, Delhi, Chennai, Hyderabad, and Bangalore), b) other urban areas. Similarly, rural areas are divided into two groups based on an index of infrastructural development. Villages that have ample infrastructure (such as paved roads, access to urban centers, postal and telephone connections, access to electricity that powers lights and television) are coded as more developed villages; and villages that have poor infrastructure are coded as less developed villages.

9. Descriptive Statistics

Two-child norm seems to have taken a firm hold in India. In the overall national sample, about 1 per cent women said they didn't want any children. Figure 2 shows that 4.1% had an ideal family size of 1, 58.9% women had an ideal family size of 2, 23.1% women had an ideal family size of 3, and about 12.8% had an ideal family size of 4 or more.

Figure 2 Distribution of ideal family size in 2005 (weighted)



Sample size, n=19,132, using IHDS-1 data

Table 2 shows the distribution of ideal family size across different states. This table also shows a strong preference for two-child families in most states. The two States that are exceptions with respect to this are Bihar and Uttar Pradesh. In Bihar, majority women had an ideal family size of 4 or more children (36.29%), followed by an ideal family size of 3 children for 36.29%; whereas in Uttar Pradesh, majority women (35.42 %) had an ideal family size of 3, and the second highest preferred family size was 2 children (for 33.92% women).

Table 2 Distribution of state-wise ideal family size in 2005# (weighted)

| State | 0 | 1 | 2 | 3 | 4+ | Missing |
|-------------------|-------------|------------|--------------|--------------|--------------|-------------|
| All India | 1.04 | 3.9 | 55.53 | 21.84 | 12.01 | 5.68 |
| Jammu and Kashmir | 2.31 | 1.36 | 37.7 | 35.32 | 13.16 | 10.14 |
| Himachal Pradesh | 0.92 | 15.39 | 76.68 | 4.75 | 1.29 | 0.97 |
| Uttarakhand | 0 | 3.58 | 53.35 | 25.55 | 15.16 | 2.36 |
| Punjab | 12.43 | 4.54 | 61.07 | 15.03 | 4.21 | 2.72 |
| Haryana | 0.73 | 1.85 | 61.61 | 23.49 | 1.66 | 10.65 |
| Delhi | 0.41 | 0.38 | 71.17 | 17.43 | 6.49 | 4.12 |
| Uttar Pradesh | 1.23 | 1.56 | 33.92 | 35.42 | 21.57 | 6.29 |
| Bihar | 0 | 2.79 | 22.05 | 35.93 | 36.29 | 2.93 |
| Jharkhand | 0 | 1.35 | 43.53 | 35.1 | 17.07 | 2.94 |
| Rajasthan | 0.39 | 1.91 | 61.99 | 22.51 | 6.7 | 6.51 |
| Chhattisgarh | 0 | 0.55 | 53.03 | 32.68 | 8.38 | 5.36 |
| Madhya Pradesh | 0.06 | 0.96 | 57.99 | 20.88 | 16.85 | 3.25 |
| Northeast | 1.43 | 9.65 | 42.37 | 9.11 | 22.66 | 14.78 |
| Assam | 2.22 | 0.61 | 39.65 | 26.22 | 11.59 | 19.71 |
| West Bengal | 0.63 | 11.59 | 68.54 | 9.35 | 3.61 | 6.28 |

| State | 0 | 1 | 2 | 3 | 4+ | Missing |
|-------------------|------|------|-------|-------|-------|---------|
| Orissa | 0 | 4.71 | 62.2 | 21.01 | 8.2 | 3.88 |
| Gujarat | 0.24 | 4.27 | 72.22 | 14.97 | 0.77 | 7.54 |
| Maharashtra & Goa | 0.52 | 2.59 | 61.5 | 17.44 | 10.62 | 7.33 |
| Andhra Pradesh | 1.71 | 4.03 | 65.35 | 19.71 | 4.87 | 4.34 |
| Karnataka | 3.61 | 3.68 | 67.15 | 10.11 | 8.79 | 6.66 |
| Kerala | 0.29 | 3.48 | 71.26 | 14.1 | 4.65 | 6.21 |
| Tamil Nadu | 0.5 | 8.98 | 75.75 | 8.32 | 3.68 | 2.77 |

#Indicates proportion with ideal family size 0,1,2,3,4 and more, across States.
n=20,464

Table 3 shows average ideal family size in 2005 (column 1), and the percentage women, who had at least one unintended birth between 2005, and 2012 (column 2) across 22 state groups. The average ideal family size in 2005 in India was 2.45, the average in EAG states was higher than the all India average at 2.74, while the average ideal family size for non-EAG states was lower (2.21) than the all India average. While looking separately at the 22 States, it can be seen that average ideal family size in 2005, was greater than the all India average in the States of Jammu and Kashmir, Uttarakhand, Uttar Pradesh, Bihar, Jharkhand, Chhattisgarh, Madhya Pradesh, Assam and other Northeastern States. There was only one state: Bihar (3.18), where average ideal family size in 2005 was above 3 children.

Table 3 Ideal family size in 2005 and Unintended Births between 2005 and 2012*

| State | Average Ideal Family size (1) | Percentage women who have at least one unintended birth between 2005 and 2012+ (2) | TFR 2005-06 (3) |
|-------------------|----------------------------------|---|--------------------|
| All India | 2.45 | 17.56 | 2.7 |
| EAG States | 2.74 | 25.42 | 3.1 |
| Non-EAG States | 2.21 | 11.7 | 2.4 |
| Jammu and Kashmir | 2.64 | 14.8 | 2.4 |
| Himachal Pradesh | 1.9 | 23.33 | 1.9 |
| Uttarakhand | 2.57 | 36.61 | 2.6 |
| Punjab | 1.95 | 15.39 | 2 |
| Haryana | 2.26 | 22.97 | 2.7 |
| Delhi | 2.31 | 29.28 | 2.1 |
| Uttar Pradesh | 2.85 | 23.99 | 3.8 |
| Bihar | 3.18 | 33.51 | 4 |
| Jharkhand | 2.74 | 25.42 | 3.3 |
| Rajasthan | 2.37 | 23.98 | 3.2 |
| Chhattisgarh | 2.54 | 20.74 | 2.6 |
| Madhya Pradesh | 2.57 | 23.03 | 3.1 |
| Northeast | 2.67 | 15.05 | 3.1 |

| State | Average Ideal Family size | Percentage women who have at least one unintended birth between 2005 and 2012+ | TFR 2005-06 |
|-------------------|---------------------------|--|-------------|
| | (1) | (2) | (3) |
| Assam | 2.57 | 13.78 | 2.4 |
| West Bengal | 2.04 | 14.76 | 2.3 |
| Orissa | 2.34 | 17.66 | 2.4 |
| Gujarat | 2.13 | 23.38 | 2.4 |
| Maharashtra & Goa | 2.39 | 10.74 | 2 |
| Andhra Pradesh | 2.24 | 8.57 | 1.8 |
| Karnataka | 2.2 | 9.73 | 2.1 |
| Kerala | 2.22 | 3.75 | 1.9 |
| Tamil Nadu | 2.06 | 4.5 | 1.8 |
| Sample Size | 19,132 | 13,128 | |

* Calculations for columns (1) and (2) based on data for ever-married women aged 18-40 from IHDS 2005, and 2012. TFRs reported in column (3) come from National Family Health Survey III, 2005-06 report for women aged 15-49. Average TFR for EAG and non-EAG States based on State-groups used from NFHS.

Column 2 in Table 3 shows the state-wise distribution of the percentage women who had at least one unintended birth between 2005 and 2012. As mentioned earlier, this is limited to a smaller sample size of 13,128 women, who had children less than or equal to ideal family size in 2005. We find that almost 18% of women in this sample had at least one unintended birth between 2005 and 2012, i.e. exceeded ideal family size as expressed in Wave 1 interview. The percentage of women who had unintended births was much higher in EAG States (more than 25%), and lower in non-EAG States (almost 12%) compared to the national average. Proportion of women who had at least one unintended births between 2005 and 2012 were the lowest in Kerala (less than 4%), followed by Tamil Nadu (less than 5%), Andhra Pradesh (less than 9%), and Karnataka (less than 10%); and the highest in Uttarakhand (36.6%), followed by Bihar (33.5%), Delhi (29.3%), Jharkhand (25.4%), and Uttar Pradesh (24%). It is important to remember, that the IHDS is not designed to be representative at the state level and small samples render state-wise ranking indicative rather than definitive. Column 3 in Table 3 shows the total fertility rates (TFR) for women aged 15-49 aggregated from National Family Survey (NFHS III, 2005-06) report, and as depicted in figure 1.

Summary statistics for the independent variables in this analysis are presented in Appendix Table 1. We present the distribution of our two primary fertility outcomes, desired fertility, and unplanned fertility across different independent variables in Table 4. Bivariate statistics show a wide variation in fertility intentions (column 1), and behavior (column 2) across education levels, income, area of residence, number of decisions in which a woman has the most say, exposure to mass media, any connection to formal social networks, and maternal health care utilization during the last birth, as reported in 2005.

Table 4 Distribution of indicators of Fertility intentions and behavior by key independent variables (%)

| | Ideal Family size in 2005 (1) | Has an unintended birth between 2005 and 2012+ (2) |
|--|--------------------------------------|---|
| Education in 2005 | | |
| Illiterate | 2.69 | 20.95 |
| Pre Primary | 2.45 | 16.61 |
| Primary | 2.27 | 17.41 |
| Secondary | 2.1 | 10.83 |
| High Secondary | 2.07 | 11.21 |
| College | 1.92 | 8.96 |
| Asset Quintile | | |
| Poorest | 2.82 | 24.22 |
| Second Quintile | 2.66 | 21.72 |
| Third Quintile | 2.4 | 17.8 |
| Fourth Quintile | 2.25 | 13.57 |
| Richest | 2.1 | 10.91 |
| Area of Residence | | |
| Metro Cities | 2.07 | 10.81 |
| Other Urban | 2.25 | 14.22 |
| More Developed Villages | 2.4 | 14.52 |
| Less Developed Village | 2.63 | 22.7 |
| Number of decisions in the household where the woman has the most say | | |
| 0 | 2.48 | 18.84 |
| 1 | 2.45 | 17.28 |
| 2 | 2.28 | 12.69 |
| 3 | 2.38 | 15.17 |
| 4 | 2.33 | 8 |
| Number of Social Networks in the Formal Sector | | |
| 0 | 2.47 | 18.67 |
| 1 | 2.47 | 16.26 |
| 2 | 2.47 | 18.7 |
| 3 | 2.34 | 15.09 |
| Score on Exposure to Mass media | | |
| 0 | 2.73 | 23.95 |
| 1 | 2.45 | 16.95 |
| 2 | 2.34 | 16.42 |
| 3 | 2.17 | 11.12 |
| Any ANC during the last birth# | | |
| Yes | 2.35 | 23.89 |
| No | 3.04 | 39.11 |

| | Ideal Family size in 2005 (1) | Has an unintended birth between 2005 and 2012+ (2) |
|---|--|---|
| Postnatal Checkup during last birth# | | |
| No Postnatal check-up | 2.66 | 29.83 |
| First check-up after 24 hours of birth | 2.26 | 21.92 |
| First check-up within 24 hours of birth | 2.29 | 23.93 |

Limited to those who have a birth between Jan 2000-interview date in 2005

+ Limited to those who had children less than, or equal to ideal family size in 2005.

It can be seen that ideal family size in 2005 (column 1) decreases with an increase in education, and is the highest for illiterate women (about 2.7), and the lowest for college educated women (below 2). Column (2) shows that while only about 9% women with college education have at least one unintended birth, almost 21% women who are illiterate have at least one unintended birth. A similar pattern is observed by household socio-economic status. The ideal family size decreases with an increase in household assets, while for women belonging to poorest families ideal family size is 2.8, for women belonging to richest families it is 2.1. Percentage women having an unintended birth decrease significantly as household wealth increases. While more than 24% women belonging to poorest households have at least one unintended birth between 2005 and 2012; less than 11 % women belonging to the richest households have at least one unintended birth.

Differences in ideal family size across areas of residence, show that ideal family size is highest in less developed villages (2.6), followed by more developed villages (2.4), other urban areas (2.3), and lowest in metro cities (2.1). Percentage women having at least one unintended birth, is the highest in less developed villages (almost 23%), and the lowest in metro cities (less than 11%).

Next we look at the relationship between the number of decisions in which women relatively have the most say, and fertility intentions and behavior. Women who do not have the most say in any household decisions have the highest average ideal family size in 2005, compared to women who have the most say in all 4 major household decisions. A greater percentage of women who didn't have the most say in any household decision were likely to have an unintended birth between 2005 and 2012, compared to women who had the most say in all 4 household decisions.

Both ideal family size in 2005, and percentage women who had unintended births between 2005 and 2012 did not vary much by the number of social networks in the formal sector; though those who had connections with all three social networks had a lower ideal family size; and also a smaller percentage of women with all three social networks had unintended births, compared to those with no connections in the formal sector.

Ideal family size in 2005 is higher on an average for women who reside in households that have exposure to fewer forms of mass media. Percentage women having unintended births between 2005 and 2012 is also higher if women reside in households that have no exposure to mass media (about 24%), compared to women who reside in households where women have exposure to all forms of mass media (about 11%).

Finally, variables measuring exposure to maternal healthcare utilization during the last birth are indicative of connectivity with the maternal and child health system, and thus the availability of healthcare facilities in the region of residence. Women who had obtained at least one ANC during her last birth (as reported in 2005), was likely to have a lower average ideal family size in 2005 (2.4), compared to women who received no ANC (3) during her last birth. A larger percentage of women who didn't receive any ANC during their last birth were likely to have at least one unintended birth (more than 39%), compared to those who received at least one checkup (about 24%). Similarly, for PNC use during the last birth (as reported in 2005), average ideal family size in 2005 was lower for women who received any postnatal checkup, compared to those who did not. Also, a higher percentage of women didn't obtain any postnatal checkup during their last birth (as reported in 2005) were likely to have at least one unintended birth between 2005 and 2012.

10. Results from Multi-level Regressions

Results from hierarchical multi-level regressions are presented in Tables 5 and 6. Each table contains three columns. Column 1 includes only state-level variation at the second level without any independent variables; column 2 adds individual level correlates of fertility preferences, and behaviors; and column 3 adds contact with health services for antenatal and postnatal care.

As the discussion of statistical models in section 7 notes, this approach allows us to partition the variance into two components, within and between states. The Intraclass Correlation Coefficient describes the component of variance that can be attributed to between state variance. In as much as between state differences are a function of individual characteristics such as education and economic status, addition of these factors should reduce the size of ICC between models 1, 2 and 3.

10.1. *Desired Family Size.*

Table 5 shows the results from the hierarchical linear regression model for the family size women would have wanted had they been able to choose their family at the start of their reproductive career. This variable was reported at Wave 1 interview in 2005. Model 1 is the basic model without any individual/household level predictors. This model enables us to obtain the baseline ICC, with the goal of decreasing it's magnitude upon adding other individual and household variables (in model 2), and variables indicative of maternal healthcare utilization (model 3). ICC of 0.13 in model 1 suggests that 13% of the variation in ideal family size in 2005 is between states, and 87% of the variation in ideal family size in 2005 is between women residing in the same state. Model 2 adds the individual, and household level independent, and other control variables to model 1. Note that the within state component includes both systematic variation between women based on their education, economic status, social networks and a host of other individual characteristics identified in the literature reviewed above as determinants of fertility preferences as well an unmeasured component reflecting idiosyncratic variation between individuals. Adding individual /household level factors reduces the ICC from about 0.13 to 0.09, which indicates a 35% decline between model 1 and 2. This means that a substantial part of the inter-state variation in ideal family size is explained by the variables included in model 2. In

Model 3, we add contact with the health system during prior delivery. Addition of these variables reduces the ICC to 0.08, which indicates a 39% decline from the null model. Thus, individual and household level factors; and connection that women have with maternal and child health workers explains a significant part of inter-state variation in ideal family size.

Table 5 Beta coefficients from hierarchical linear regression examining women's ideal number of children in 2005

| | Model 1 | Model 2 | Model 3 |
|---|---------|-----------------------|-----------------------|
| Education (Reference group: Illiterate) | | | |
| Incomplete primary | | -0.037 (0.0341) | -0.030 (0.0341) |
| Primary | | -0.124*** (0.0226) | -0.113*** (0.0226) |
| Secondary | | -0.162*** (0.0347) | -0.149*** (0.0347) |
| Higher secondary | | -0.135** (0.0440) | -0.119** (0.0440) |
| College and higher | | -0.179*** (0.0498) | -0.166*** (0.0498) |
| Asset Quintile(Reference group: Poorest) | | | |
| Second quintile | | -0.110*** (0.0268) | -0.103*** (0.0268) |
| Third quintile | | -0.186*** (0.0293) | -0.172*** (0.0293) |
| Fourth quintile | | -0.241*** (0.0343) | -0.222*** (0.0344) |
| Richest | | -0.231*** (0.0416) | -0.208*** (0.0417) |
| Area of Residence(Reference group: Metro Cities) | | | |
| Other Urban | | 0.019 (0.0470) | 0.029 (0.0469) |
| More Developed Villages | | 0.059 (0.0475) | 0.064 (0.0474) |
| Less Developed Village | | 0.134** (0.0481) | 0.135** (0.0480) |
| Religion(Reference group:Hindu) | | | |
| Muslim | | 0.353*** (0.0274) | 0.356*** (0.0274) |
| Other Religion | | 0.114** (0.0395) | 0.120** (0.0395) |
| Caste Group(Reference group: Forward Caste Groups) | | | |
| Scheduled Castes (SC) | | 0.036 (0.0260) | 0.036 (0.0260) |
| Scheduled Tribes (ST) | | 0.253*** (0.0368) | 0.241*** (0.0368) |
| Other Backward Classes (OBC) | | 0.067** (0.0223) | 0.065** (0.0222) |

| | Model 1 | Model 2 | Model 3 |
|---|----------------------|-----------------------|-----------------------|
| Number of living children | | 0.218*** (0.0102) | 0.213*** (0.0102) |
| Age Category(Reference group: 18-20) | | | |
| 21-25 | | -0.084* (0.0328) | -0.083* (0.0328) |
| 26-30 | | -0.122*** (0.0363) | -0.120*** (0.0363) |
| 31-36 | | -0.045 (0.0383) | -0.044 (0.0382) |
| 36-40 | | 0.021 (0.0476) | 0.016 (0.0476) |
| Number of household decisions where the woman has most say (Reference Group: no decisions) | | | |
| 1 | | -0.046* (0.0204) | -0.045* (0.0203) |
| 2 | | -0.092** (0.0324) | -0.090** (0.0324) |
| 3 | | -0.094 (0.0600) | -0.090 (0.0599) |
| 4 | | -0.045 (0.0617) | -0.027 (0.0617) |
| Number of formal social networks(Reference Group: no connection) | | | |
| 1 | | 0.033 (0.0221) | 0.035 (0.0221) |
| 2 | | 0.005 (0.0252) | 0.005 (0.0252) |
| 3 | | -0.077** (0.0272) | -0.076** (0.0272) |
| Exposure of women in to household to mass media (Reference Group: no exposure) | | | |
| 1 | | -0.002 (0.0234) | 0.002 (0.0233) |
| 2 | | -0.013 (0.0265) | -0.009 (0.0265) |
| 3 | | 0.004 (0.0353) | 0.009 (0.0353) |
| Any antenatal check-up | | | -0.111*** (0.0236) |
| Postnatal Checkup (Reference Group: No post natal checkup) | | | |
| First check-up after 24 hours of birth | | | -0.028 (0.0221) |
| First check-up within 24 hours of birth | | | -0.075** (0.0267) |
| Constant | 2.411*** (0.0694) | 2.047*** (0.0807) | 2.136*** (0.0814) |

| | Model 1 | Model 2 | Model 3 |
|--|-----------------------|-----------------------|-----------------------|
| Random effects Parameters | | | |
| Constant (level 2 State of Residence)) | -1.137*** (0.1540) | -1.485*** (0.1572) | -1.517*** (0.1580) |
| Constant (residual) | -0.191*** (0.0077) | -0.297*** (0.0077) | -0.299*** (0.0077) |
| ICC | 0.131 | 0.085 | 0.08 |
| Observations | 8,348 | 8,348 | 8,348 |
| Number of State Groups | 22 | 22 | 22 |

Standard errors in parentheses

*** p<0.001, ** p<0.01, * p<0.05, + p<0.10

Results from model 2 show that, of key socio-economic and demographic variables; being illiterate (compared to having primary or higher level education); belonging to poorest households (compared to that for richer households); belonging to Muslim and other religious groups (compared to Hindu groups); residing in a less developed village (compared to a metro city); belonging to ST and OBC caste groups (compared to forward caste groups); having a higher number of living children; and being 18-20 (compared to being 21-30 is associated with having higher ideal family size in 2005. Having three formal social networks (compared to no social networks in the formal sector), and, having the most say in more number of household decisions (having the most say in one or two decisions compared to not having the most say in any decision) is linked to having smaller ideal family size. Finally, in model 3 ANC, and PNC use in 2005 are added to model 2, while the significance level and directions for all other independent and control variables remain similar to that in model 2, model 3 shows that obtaining at least one ANC (compared to obtaining none), and the first PNC within 24 hours of birth (compared to obtaining none), during the last birth is associated with having a smaller ideal family size.

10.2. Unplanned Birth

Table 6 shows the results from the hierarchical logistic regression model examining the likelihoods of having at least one unintended birth between 2005 and 2012. Model 1 is the basic model without any individual/household level predictors. ICC of about 0.10 in this model.⁷ Adding individual, and household level variables in model 2, reduces the variance of the random effects coefficient and reduced the ICC from 0.101 to 0.095 i.e. about 6% decline between model 1 and model 2. Adding the ANC, and PNC use variables in model 3, reduces the variance of the random effects coefficients and reduced the ICC to 0.092, a 9% decline from the null model. Comparison of changes in ICC between the two dependent variables, desired fertility and unplanned births is striking. Whereas a vast proportion of inter-state variation in desired fertility appears to be related to differing composition of women living in these states, the same cannot

⁷ Interpretation of ICC for logistic regressions is not strictly comparable to the one for multilevel linear models continuous since it depends on the prevalence of outcome in clusters of interest. Furthermore, the between-cluster variance is defined on a different scale (e.g., the log-odds scale in a logistic regression), than the binary response scale (Austin and Merlo 2017).

be said of unintended fertility. This is more evident as we look at the variation in unintended fertility between women with different characteristics.

Table 6 Beta coefficients from hierarchical logistic regression examining likelihoods of a woman having at least one unwanted birth⁺⁺

| | Model 1 | Model 2 | Model 3 |
|--|---------|---------------------------|---------------------------|
| Education (Reference group: Illiterate) | | | |
| Incomplete primary | | -0.124 (0.1303) | -0.117 (0.1304) |
| Primary | | -0.127 (0.0844) | -0.109 (0.0848) |
| Secondary | | -0.171 (0.1327) | -0.158 (0.1331) |
| Higher secondary | | -0.258 (0.1731) | -0.239 (0.1737) |
| College and higher | | -0.465* (0.2085) | -0.448* (0.2089) |
| Asset Quintile (Reference group: Poorest) | | | |
| Second quintile | | -0.074 (0.0999) | -0.066 (0.1001) |
| Third quintile | | -0.199+ (0.1109) | -0.186+ (0.1117) |
| Fourth quintile | | -0.376** (0.1321) | -0.354** (0.1331) |
| Richest | | - 0.576*** (0.1637) | - 0.554*** (0.1648) |
| Area of Residence (Reference group: Metro Cities) | | | |
| Other Urban | | -0.152 (0.1956) | -0.138 (0.1958) |
| More Developed Villages | | -0.343+ (0.1969) | -0.335+ (0.1970) |
| Less Developed Village | | -0.239 (0.1981) | -0.235 (0.1982) |
| Religion (Reference group: Hindu) | | | |
| Muslim | | 0.676*** (0.1044) | 0.682*** (0.1045) |
| Other Religion | | 0.127 (0.1504) | 0.130 (0.1503) |
| Caste Group (Reference group: Forward Caste Groups) | | | |
| Scheduled Castes (SC) | | 0.371*** (0.0990) | 0.370*** (0.0991) |
| Scheduled Tribes (ST) | | 0.309* (0.1504) | 0.297* (0.1503) |

| | Model 1 | Model 2 | Model 3 |
|---|---------|--------------------------------|--------------------------------|
| Other Backward Classes (OBC) | | (0.1387) -0.008 (0.0873) | (0.1391) -0.012 (0.0874) |
| Number of living children | | 0.090* (0.0436) | 0.083+ (0.0439) |
| Age Category(Reference group: 18-20) | | | |
| 21-25 | | -0.040 (0.1029) | -0.037 (0.1030) |
| 26-30 | | -0.339** (0.1215) | -0.334** (0.1216) |
| 31-36 | | - 0.794*** (0.1373) | - 0.795*** (0.1375) |
| 36-40 | | - 1.180*** (0.2310) | - 1.188*** (0.2313) |
| Number of household decisions where the woman has most say (Reference Group: no decisions) | | | |
| 1 | | -0.056 (0.0791) | -0.050 (0.0792) |
| 2 | | -0.325* (0.1387) | -0.315* (0.1388) |
| 3 | | 0.015 (0.2423) | 0.019 (0.2429) |
| 4 | | -0.011 (0.2606) | 0.018 (0.2609) |
| Number of formal social networks(Reference Group: no connection) | | | |
| 1 | | -0.198* (0.0872) | -0.196* (0.0873) |
| 2 | | 0.041 (0.0977) | 0.050 (0.0978) |
| 3 | | 0.075 (0.1055) | 0.089 (0.1058) |
| Exposure of women in to household to mass media (Reference Group: no exposure) | | | |
| 1 | | 0.015 (0.0879) | 0.020 (0.0880) |
| 2 | | -0.004 (0.0998) | 0.002 (0.1000) |
| 3 | | -0.057 (0.1389) | -0.053 (0.1390) |
| Any antenatal check-up | | | -0.206* |

| | Model 1 | Model 2 | Model 3 |
|---|-----------------------|---------------------|---------------------|
| | | | (0.0884) |
| Postnatal Checkup (Reference Group: No post natal checkup) | | | |
| First check-up after 24 hours of birth | | | -0.030 (0.0868) |
| First check-up within 24 hours of birth | | | 0.109 (0.1024) |
| Constant | -1.155*** (0.1354) | -0.613* (0.2789) | -0.481+ (0.2851) |
| Random effects Parameters | | | |
| Var(_ Constant (level 2 State of Residence)) | 0.369** (0.1258) | 0.345** (0.1216) | 0.332** (0.1187) |
| ICC | 0.101 | 0.095 | 0.092 |
| Observations | 5,903 | 5,903 | 5,903 |
| Number of State Groups | 22 | 22 | 22 |

Standard errors in parentheses

*** p<0.001, ** p<0.01, * p<0.05, + p<0.10

++ Limited to those who had children less than or equal to ideal family size in 2005, and those who had at least one birth between Jan 2000 and interview date in 2005. The dependent variable measure whether these women exceed their ideal family size in 2005, and go on to have at least one unwanted birth between 2005 and 2012.

Results from model 2 show that while education, and income reduce the likelihood of having unplanned births, women need to reach very high levels of education or income before this effect is large enough to be significant. Whereas for desired fertility, even a small amount of education has a large and significant effect, for unplanned births, college education is required before we see any impact. Similarly, there is only a small reduction in likelihood of unplanned births between women in the bottom 60 percent of household wealth index.

Age, however, appears to have an interesting relationship with unplanned births. Younger women seem to be particularly susceptible to experiencing unplanned birth. Since they are also most likely to desire smaller families, this is clearly a niche where unmet need for contraception appears to be the greatest. Unplanned births appear to be higher among Muslim, Scheduled Caste and Scheduled Tribe households, identifying another population that seem to be poorly served by existing service delivery infrastructure. Among other factors we control for, having no connections in the formal sector (compared to having one connection), and having the most say in no household decisions (compared to having the most say in a couple of household decisions) is associated with greater likelihoods of at least one unintended birth between 2005 and 2012. Results from model 3 add variables indicating a woman's connection with the maternal and child health system. Results from model 3 show that obtaining at least one antenatal checkup during their last birth (as reported in 2005) is associated with lower likelihoods of unintended birth between 2005 and 2012.

11. Discussion and Conclusion

This chapter has examined regional differences in fertility in India through the prism of desired and unintended fertility using prospective data. Descriptive statistics presented in Table 3 show tremendous variation between states. Whereas Southern state of Karnataka had desired family size of 2.2 children with less than 10% women having an unintended birth, Bihar in Northern India had desired fertility of 3.18 with about a third of the women reporting unintended birth. While these results highlight well established regional differences in fertility with states in the Hindi speaking heartland of India showing greatest desired as well as unintended fertility, they also shed interesting light on possible processes through which these differences emerge.

States in India are almost like mini nations, they differ tremendously in social, political and economic culture and have tremendous independence in running their health programs. Even when central government pays for some the major schemes such as maternal protection schemes, its operation is left to states creating very different service delivery climate with some states able to efficiently administer these schemes and others being lackadaisical in their governance.

Our results suggest that a substantial proportion of variation in fertility preferences between states can be explained by individual factors such as education, and income. Poorer states with less educated populations such as, Bihar and Uttar Pradesh have higher desired family size. Almost 40 percent of inter-state variation in fertility preferences is reduced once we take into account differences in these individual attributes. However, this same, well established pathway to reduced fertility does not apply easily to unintended fertility. Education, and income reduce unintended births, but only after they reach a very high threshold. Whereas, even secondary enrolment reduces desired family size, it requires college education to make a significant dent in the likelihood of unintended births. This lack of association at individual level translates into the fact that between state variance in unintended births remains impervious to addition of individual variables with only 9 percent decline in between-state ICC.

This suggests that the unobserved component of interstate variance is very strong, and not easily explainable through individual factors. Some of this could be attributed to the exclusion of socio-cultural differences across regions such as differences in kinship patterns, extent of coresidence with parents, and gender relations (Desai et al. 2010; Dyson and Moore 1983). It is also possible that regional differences in the quality of family planning services may result in large inter-state variation. Although we tried to capture state provision of health services through variables that capture use of antenatal and postnatal care, addition of these variables do not seem to make a dent in interstate variance. We argue that this is because the measures of service delivery we have access to are focused on maternal and child health services and may be quite distinct from provision of family planning services. While the same health workers are supposed to provide maternal and child health services as well as family planning services, these are two distinct activities and it is not possible to assume that the providers who prioritize one activity will also prioritize the other.

Our results also highlight two types of populations that seem to most at risk of having unintended fertility. Young women below 25 have a preference for having smaller families, but they also seem to be most at risk of being unable to implement their preferences. Muslims, Scheduled Castes, and Scheduled Tribes also seem to be most at risk of having unintended

births. These observations identify niche populations that may be most at risk of being excluded from the service delivery systems may require special programming.

At a broader level, these results document an interesting puzzle. Individual factors that have been long established in fertility research, education, economic status, gender relations in the household and social networks, seem to play an important role in shaping fertility preferences. However, our ability to predict unintended fertility seems to be far more limited. Even education, the greatest weapon in the armory of demographic research, seems to be only weakly correlated with unintended births. This should be an important area of future research.

It is important to note possible limitations of this analysis as we interpret the results. Firstly, for measuring fertility preferences we use 'ideal family size' as the key dependent variable, this measure could be subject to ex-post rationalization, whereby women adjust their ideal family size to be near their actual family size, which causes an overestimation of wanted fertility (for e.g. Bongaarts, 1990,2011; Bhushan and Hill, 1995; Lightbourne, 1985; Westoff, 1991). However, by limiting our sample to those who had children lower than, or equal to ideal family size in 2005; and subsequently checking whether this sub-group of women had more number of living children in 2012, than their ideal family size in 2005, enables us to partially address this limitation. Secondly, we have assumed that fertility preferences as expressed at Wave 1 interview reflect long term predisposition on the part of individuals towards having a particular number of children and taken deviation from over the subsequent eight years as unintended fertility. However, individuals may have adjusted their fertility preferences upwards, and what we define as unintended fertility may well reflect a change in preference rather than inability to implement preferences. In spite of this caveat, we suggest that longitudinal data on preferences and their implementation provide us with a way of looking at regional differences in fertility that has not been feasible to date with cross-sectional data and deserves further attention.

In conclusion, our study uses data from the first nationally representative panel survey in India, and examines factors that impact regional differences in fertility preferences, and on the ability to crystallize these preferences. Results highlight that while socio-economic traits of individuals and their household's account for a significant portion of regional differentials in fertility preferences, they account for a much smaller proportion of differentials in unintended births across regions. Our study implies that unobserved factors such as cultural norms, social structure, and particularly the quality of regional healthcare systems could play a much bigger part in explaining differences in unintended births compared to fertility preferences.

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Appendix Tables

Appendix Table 1: Weighted summary statistics for analytical sample of women in India aged 18–40 in 2005

| | N | Percentage/Mean (Standard Deviation) |
|--|--------|--|
| <i>Indicators of fertility desires and unintended births</i> | | |
| Ideal number of children desired in 2005 | 19,132 | 2.448(0.881) |
| Percentage women who had unintended births between 2005 and 2012+ | 13,128 | 17.56 |
| <i>Independent Variables</i> | | |
| Score on Exposure to Mass media | | |
| 0 | 19,132 | 24.67 |
| 1 | 19,132 | 32.21 |
| 2 | 19,132 | 29.39 |
| 3 | 19,132 | 13.73 |
| Number of formal social networks | | |
| 0 | 19,132 | 45.15 |
| 1 | 19,132 | 21.81 |
| 2 | 19,132 | 16.69 |
| 3 | 19,132 | 16.35 |
| Number of household decisions (except cooking) where the woman has most say | | |
| 0 | 19,132 | 64.12 |
| 1 | 19,132 | 22 |
| 2 | 19,132 | 8.12 |
| 3 | 19,132 | 2.26 |
| 4 | 19,132 | 3.49 |
| Postnatal Checkup | | |
| None | 8,453 | 67.05 |
| First check-up after 24 hours of birth | 8,453 | 20.46 |
| First check-up within 24 hours of birth | 8,453 | 12.49 |
| Any antenatal check-up | 8,408 | 74.59 |
| Asset Quintile | | |
| Poorest | 19,132 | 18.20 |
| Second quintile | 19,132 | 20.10 |
| Third quintile | 19,132 | 24.29 |
| Fourth quintile | 19,132 | 21.05 |
| Richest | 19,132 | 16.26 |
| Number of living children in 2005 | 19,132 | 2.3(1.196) |
| Area of residence | | |
| Metro Urban | 19,132 | 5.48 |
| Other Urban | 19,132 | 17.91 |
| More Developed Villages | 19,132 | 35.87 |

| | N | Percentage/Mean (Standard Deviation) |
|------------------------------|--------|--|
| Less Developed Villages | 19,132 | 40.73 |
| Religion | | |
| Hindu | 19,132 | 83.73 |
| Muslim | 19,132 | 11.23 |
| Other Religion | 19,132 | 5.04 |
| Caste Group | | |
| Forward Castes (FC) | 19,132 | 26.80 |
| Scheduled Castes (SC) | 19,132 | 23.10 |
| Scheduled Tribes (ST) | 19,132 | 7.17 |
| Other Backward Classes (OBC) | 19,132 | 42.93 |
| Age Category | | |
| 18-20 | 19,132 | 8.31 |
| 21-25 | 19,132 | 21.87 |
| 26-30 | 19,132 | 16.16 |
| 31-36 | 19,132 | 31.03 |
| 36-40 | 19,132 | 22.63 |
| Education | | |
| Illiterate | 19,132 | 47.85 |
| Incomplete primary | 19,132 | 7.73 |
| Primary | 19,132 | 28.67 |
| Secondary | 19,132 | 8.54 |
| Higher secondary | 19,132 | 4.04 |
| College and higher | 19,132 | 3.17 |

+ Limited to those who had children less than or equal to ideal family size in 2005.